Estimating the anomalous diffusion exponent for single particle tracking data with measurement errors. FIMA approach

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KIAS Workshop on Anomalous Dynamics in Biological Systems, 2015

Krzysztof Burnecki Estimating the anomalous diffusion exponent

Outline

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• Autoregressive fractionally integrated moving average (ARFIMA) model

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- Single particle tracking data with measurement errors. MSD vs. FIMA approach (Burnecki, Kepten, Garini, Sikora and Weron, Scientific Reports, 2015)



ARFIMA (another acronym FARIMA) process

Autoregressive fractionally integrated moving average (ARFIMA) process $\{X_t\}$, denoted by ARFIMA(p, d, q), introduced by Granger and Joyeux (J. Time Ser. Anal., 1980) and Hosking (Biometrika, 1981), is defined by equation

$$(1-B)^d [X(t) - \phi_1 X(t-1) - \ldots - \phi_p X(t-p)] = Z(t) - \psi_1 Z(t-1) - \ldots - \psi_q Z(t-q),$$

where

- B is the shift operator, i.e. $B^{j}X_{t} = X_{t-j}$,
- Z(t) is the noise sequence with either finite or infinite variance,
- $\Phi_p(B) = 1 \phi_1 B \phi_2 B^2 \ldots \phi_p B^p$ is the autoregressive (AR) polynomial and
- Ψ_q(B) = 1 ψ₁B ψ₂B² ... b_qB^q is the moving average (MA) polynomial.

What is new in the definition is the fractional difference operator $(1-B)^d$ with the exponent *d*, called the memory parameter.

The operator has the infinite binomial expansion

$$(1-B)^d = \sum_{j=0}^{\infty} \binom{d}{j} (-B)^j = \sum_{j=0}^{\infty} \pi_j B^j,$$

where

$$\pi_j = rac{\Gamma(j-d)}{\Gamma(j+1)\Gamma(-d)},$$

and Γ is the gamma function.

If the noise Z(t) consists of i.i.d. random variables belonging to the domain of attraction of the Lévy α -stable law, then $d < 1 - 1/\alpha$ and $0 < \alpha \leq 2$. The noise may be either Gaussian, non-Gaussian with finite variance or it may have infinite variance.

- For finite variance case, the typical choices are exponential, log-normal, normal-inverse Gaussian and truncated stable distributions.
- For infinite variance variables one may consider symmetric and skewed Lévy stable, and Pareto distributions.

- It's one of the most popular examples of long memory (long-range dependent) processes.
- In the Gaussian case $(\alpha = 2)$ the autocovariance function $\tau(n) = \langle X(n)X(0) \rangle \langle X(n) \rangle \langle X(0) \rangle \sim n^{2d-1}$. Moreover, for d > 0 we have $\sum_{n=0}^{\infty} |\tau(n)| = \infty$. This serves as a classical definition of long memory.
- For α < 2 the covariance does not exist and one has to replace it, e.g., with the codifference. It was proved that for d > 1 - 2/α ARFIMA possesses long-term dependence in the classical sense.

ARFIMA times series have been widely studied in the literature in

- telecommunication,
- industry,
- economics,
- astronomy,
- hydrology,
- image processing,
- geoscience,
- medicine.

- Aggregated ARFIMA(0, 0, 0) ⇒ Lévy stable motion, in particular Brownian motion.
- Aggregated ARFIMA(0, d, 0) ⇒ fractional Lévy stable motion (FLSM) with H = d + 1/α, in particular fractional Brownian motion (FBM) with H = d + 1/2.
- Contrary to FBM and FLSM it can also describe different light and heavy-tailed distributions and a short (exponential-like) memory.
- It can also model a measurement error (MA part includes the information on the error).

ARFIMA trajectories



Figure 1 : Simulated trajectories of symmetric 1.8-stable ARFIMA(0, d, 0) series with the memory parameter $d \in \{-0.43, 0.03, 0.43\}$ (left panel), and their cumulative sums (right panel).

ARFIMA process for p = 0. FIMA(d, q)

FIMA(d, q):=ARFIMA(0, d, q) is represented by the fractional difference equation:

$$(1-B)^d X(t) = Z(t) - \psi_1 Z(t-1) - \ldots - \psi_q Z(t-q).$$

 In many applications FIMA(d, 1) model is sufficient to describe the data well. FIMA(d, 1) can be considered a first-order approximation of the arbitrary short memory structure (q lags) taking into account only the first lag. In this case the model reduces to:

$$(1-B)^d X(t) = Z(t) - \psi Z(t-1), \ t = 0, \pm 1, \dots$$

The estimator of the parameter vector β is defined as the vector argument $\tilde{\beta}$, for which the following function attains its minimum value:

$$g(\beta) = \int_0^\pi \frac{\widetilde{I}_N(\lambda)}{W(\lambda,\beta)} d\lambda,$$

where

$$W(\lambda, \beta) = rac{(1+2\psi\cos\lambda+\psi^2)}{(2-2\cos\lambda)^d}$$

is the spectral density of the FIMA process.

Sample MSD

Let {Z_i, i = 1,..., N} be a sample of length N. The sample (time average) MSD is defined as

$$M_N(\tau) = rac{1}{N- au} \sum_{k=1}^{N- au} (Z_{k+ au} - Z_k)^2.$$

 If the sample comes from an ARFIMA(p, d, q) process with noise belonging to the domain of attraction of the Lévy α-stable law, then for its partial sum process for large N

$$M_N(\tau) \sim \tau^{2d+1} S_{\alpha/2},$$

where \sim means similarity in distribution, and $S_{\alpha/2}$ is a totally skewed $\alpha/2$ -stable random variable (Burnecki and Weron, PRE 2010).

- $d > 0 \Rightarrow$ superdiffusion
- $d = 0 \Rightarrow$ diffusion
- $d < 0 \Rightarrow$ subdiffusion
- In general the memory parameter *d* controls the type of anomalous diffusion regardless of the underlying distribution

Estimating the anomalous diffusion exponent with measurement errors (Burnecki, Kepten, Garini, Sikora and Weron, Scientific Reports, 2015)

We consider the toy model:

$$M(t) = B_H(t) + \epsilon(t), \quad t = 1, 2, \dots, T,$$

where

- $\{B_H(t) : t = 1, 2, ..., T\}$ is the standard FBM, $E(B_H(t)^2) = t^{2H}$,
- { $\epsilon(t) : t = 1, 2, ..., T$ } is the measurement noise = white Gaussian noise, i.i.d. $\epsilon(t) \sim N(0, \sigma^2)$,
- $B_H(t)$ and $\epsilon(t)$ are independent.

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• For the toy model

$$M(t) = B_H(t) + \epsilon(t), \quad t = 1, 2, \dots, T$$
$$Cov (M(t+h), M(t)) = Cov (B_H(t+h), B_H(t)), \quad h \neq 0$$
$$E (M(t)^2) = E (B_H(t)^2) + E (\epsilon(t)^2) = t^{2H} + \sigma^2$$

• For its increments

$$\rho_{\mathcal{M}}(h) := \operatorname{Cov}\left(M(t+1) - M(t), M(t+1+h) - M(t+h)\right)$$

$$\rho_{M}(h) = \begin{cases} \rho_{H}(h) + 2\sigma^{2} & \text{if } h = 0\\ \rho_{H}(h) - \sigma^{2} & \text{if } h = 1\\ \rho_{H}(h) & \text{if } h > 1 \end{cases}, \ \rho_{H}(h) \sim H(2H-1)h^{2H-2}$$



Toy model covariance structure



Figure 2 : Covariance function of the increments of the toy model $M(t) = B_H(t) + \epsilon(t)$ and increments of the corresponding $B_H(t)$: a) with Hurst index H = 0.4 and measurement error $\sigma = 1$ (left), b) with Hurst index H = 0.7 and measurement error $\sigma = 1$ (right).

FIMA model covariance structure



Figure 3 : Covariance function of the FIMA(d, 1) model and the corresponding FI(d) a) with memory parameter d = -0.1 and $\Psi = 0.15$ (left), b) with memory parameter d = 0.2 and $\Psi = 0.6$ (right).

Our goal is to estimate from simulated trajectories of the toy model

$$M(t) = B_H(t) + \epsilon(t), \quad t = 1, 2, \dots, T,$$

the anomalous diffusion exponent $\alpha=2H$

$$E(M(t)^2) = E(B_H(t)^2) + E(\epsilon(t)^2) = t \overbrace{2H}^{\alpha} + \sigma^2$$

We apply two different estimators of α :

- time average mean square displacement (MSD)
- estimator of the memory parameter d of the FIMA process

Regimes:

- strong subdiffusion: $\alpha \in \{0.4, 0.5, 0.6\}$
- weak subdiffusion: $\alpha \in \{0.7, 0.8, 0.9\}$
- classical diffusion: $\alpha = 1$
- weak superdiffusion: $\alpha \in \{1.1, 1.2, 1.3\}$
- strong superdiffusion: $\alpha \in \{1.4, 1.5, 1.6\}$
- Measurement error $\sigma \in \{0.25, 0.5, 0.75, 1, 1.25, 1.5, 1.75, 2\}$ and the length of the trajectory T = 1024.

We simulate 1000 trajectories of the toy model for each case, estimate α from MSD and FIMA and compare the two methods.

Estimation results: strong subdiffusion



Estimation results: weak subdiffusion



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Estimation results: classical diffusion



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Estimation results: weak superdiffusion



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Estimation results: strong superdiffusion



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Estimating the anomalous diffusion exponent

- FIMA estimator is better than MSD for $\alpha \ge$ 0.7.
- For strong and weak subdiffusion both FIMA and MSD estimators are accurate only for small σ values.
- For weak and strong superdiffusion FIMA estimator is very accurate with low variance.

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Thank you!



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